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HRF estimation improves sensitivity of fMRI encoding and decoding models

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Abstract—Extracting activation patterns from functional Magnetic Resonance Images (fMRI) datasets remains challenging in rapid-event designs due to the inherent delay of blood oxygen level-dependent (BOLD) signal. The general linear model (GLM) allows to estimate the activation from a design matrix and a fixed hemodynamic response function (HRF). However, the HRF is known to vary substantially between subjects and brain regions. In this paper, we propose a model for jointly estimating the hemodynamic response function (HRF) and the activation patterns via a low-rank representation of task effects. This model is based on the linearity assumption behind the GLM and can be computed using standard gradient-based solvers. We use the activation patterns computed by our model as input data for encoding and decoding studies and report performance improvement in both settings.

Index Terms—fMRI; hemodynamic; HRF; GLM; BOLD; encoding; decoding

I. INTRODUCTION

The use of decoding models [1] to predict the cognitive state of a subject during task performance has become a popular analysis approach for fMRI studies. The converse approach is the voxel-based encoding model, which describes the information about the stimulus or task that is represented in the activity of a single voxel [2].

The input to both types of analysis consists of activation patterns corresponding to different tasks or stimulus types. These activation patterns are straightforward to calculate for blocked trials or slow-event designs, but for rapid-event designs the evoked BOLD response for adjacent trials will overlap in time, complicating the identification task.

The general linear model (GLM) was proposed [3] to overcome this difficulty. It estimates the activation patterns evoked by separate events and allows for rapid-event settings. The GLM relies on a known form of the hemodynamic response function (HRF) to estimate the activation pattern. However, it is known [4] that the shape of this response function can vary substantially across subjects and brain regions.

In this study we propose to learn the specific form of the HRF in each brain voxel to improve the computation of the activation vectors in the GLM. Joint estimation of HRF and activation patterns has already been proposed in the literature, both within the frequentist [5] and Bayesian framework [6].

We propose a model based on the linearity assumption behind GLM and low-rank factorization. We are interested in assessing the impact on higher-level analysis such as encoding and decoding studies. In particular, we examine whether encoding and decoding models give significantly different results when the activation patterns used as input data are computed using this joint estimation method.

Notation: \( \| \cdot \| \) denotes the euclidean norm for vectors. \( I \) denotes the identity matrix and \( c_i \) denotes its \( i \)th column vector. \( \otimes \) denotes the Kronecker product and \( \text{vec}(A) \) denotes the concatenation of the columns of a matrix \( A \) into a single column vector.

II. HRF ESTIMATION VIA LOW-RANK APPROXIMATION

We denote the observed fMRI time series for a single voxel by \( y = (y_1, \ldots, y_n) \) where \( y_i \) is the measurement at time \( i TR \) with \( TR \) being the time of repetition and \( n \) is the number of scans within the session. As reference HRF we use the one described by Glover [7], and denote it by \( h_c \). In this case, the GLM model specifies the observed BOLD signal as:

\[
y = X h_c \beta + P w + \varepsilon ,
\]

where \( X h_c \in \mathbb{R}^{n \times p} \) is the design matrix, i.e. the matrix whose columns are a discrete convolution of \( h_c \) with the binary stimulus vector \( V \), for condition \( i \), \( P \in \mathbb{R}^{n \times q} \) is the matrix of confounds (drifts, motions etc.) and \( \varepsilon \in \mathbb{R}^n \) is the vector of residuals, which is modeled as an auto-regressive process of order 1 (AR1(1)) to take into account the temporal correlations in the noise. We now define the binary matrix \( \hat{X} \in \mathbb{R}^{n \times pr} \) as

\[
\hat{X} = \sum_{k=0}^{r} L_k(V \otimes e_k^T)
\]

where \( r \) is the duration of the HRF in multiples of TR, \( V \) the binary stimulus matrix and \( L_k \) is the lower shift matrix of order \( k \), i.e. the matrix that shifts downwards \( k \) places all elements. Figure [1] illustrates the relationship between the event matrix and the design matrix. Matrix \( \hat{X} \) has the property that \( \hat{X} \text{vec}(h \beta^T) = X h \beta \), for any vector \( h \), allowing us to write equation (1) with an explicit linear dependency on \( h \),

\[
y = \hat{X} \text{vec}(h \beta^T) + P w + \varepsilon.
\]
The model parameters are then estimated such that they minimize the decorrelated residuals. Equivalently, we solve the following optimization problem for $\beta$, $h$ and $w$:

$$
\arg \min_{\beta, h, w} \frac{1}{2} \| Z^T (y - \hat{X} \text{vec}(h \beta^T) - P w) \|^2 ,
$$

(3)

where $Z^T$ is the whitening matrix associated with the AR(1) covariance matrix estimated by maximum likelihood. This represents a linear regression model where the coefficients are the vectorization of a rank-one matrix. It can also be seen as a Finite Impulse Response (FIR) basis GLM with a constraint such that the HRF $h$ is repeated for each condition. We will refer to this as the rank-one regression model. Note that while it is possible to add a regularization term to reflect prior knowledge of the shape of the HRF, no specific form of $h$ is privileged within this model.

A. Choice of a particular basis

In certain cases it may be desirable to reduce the number of free parameters of the HRF function within the rank-one model. An effective technique consists in constraining $h$ to be obtained as a linear combination of $t$ basis functions with $t \leq r$, in an appropriately chosen basis. Common bases used to express the HRF include the Fourier basis, polynomial basis and the canonical HRF together with its derivatives up to a certain order.

Given the subspace generated by the columns of a matrix $Q \in \mathbb{R}^{r \times t}$, we impose $h$ to lie in that subspace. In that case we can write $h = Q \alpha$ for some $\alpha \in \mathbb{R}^t$. By the properties of the Kronecker product we have the equivalence $\hat{X} \text{vec}(Q \alpha \beta^T) = X(I \otimes Q) \text{vec}(\alpha \beta^T)$. We can now define our rank-one model over parameters $\alpha$ instead of $h$ by defining the design matrix $\hat{X}_Q = \sum_{k=0}^t L_k (V \otimes q_k^T)$ where $q_i$ denotes the $i$th column of $Q$.

B. Asynchronous design

When the events are not a multiple of the repetition time, constructing the design matrix $\hat{X}$ cannot be accomplished directly by equation (2).

Let $V$ represent the binary stimulus where the events are truncated to the closest TR. Furthermore, let $V = \sum_{i=0}^m E_i$, where $E_i$ is the matrix of individual events. In the case where a continuous basis for the HRF is chosen this can be evaluated at all time. That is, the design matrix can be written as

$$
\hat{X} = \sum_{i=0}^m \sum_{k=0}^t L_k (E_i \otimes q_{k,i})
$$

where $q_{k,i}$ represents the $k$th vector in the basis evaluated at the TR timepoints plus an offset equal to the offset of event $i$ with respect to the TR.

C. Algorithm

Although the problem is not convex, the cost function is differentiable and gradient-based methods can be used to solve the optimization problem. We used the limited-memory BFGS algorithm to simultaneously optimize over parameters $\beta$, $h$, $w$. For simplicity, we have taken $Z$ to be the identity. For the general result, simply multiply $X$, $y$, $P$ by $Z^T$. Popular implementations of the algorithm only need as parameters the objective function and the gradient, given by:

$$
\nabla_\beta = (I \otimes h^T) \hat{X}^T (\hat{X} \text{vec}(h \beta^T) + P w - y)
$$

$$
\nabla_h = (\beta^T \otimes I) \hat{X}^T (\hat{X} \text{vec}(h \beta^T) + P w - y)
$$

$$
\nabla_w = P^T (\hat{X} \text{vec}(h \beta^T) + P w - y)
$$

The full gradient now can be computed by stacking $\nabla_\beta$, $\nabla_h$ and $\nabla_w$ into a single vector. Since in this algorithm only a matrix-vector product is used, the use of Kronecker product identities avoids the explicit creation of most Kronecker product matrices. We have found this implementation to take around 3 hours to perform a full brain analysis (124,000 voxels, 46 conditions) on commodity hardware with almost no increase in memory consumption once the data was loaded in memory. An implementation by the authors is publicly available [2].

III. Validation

We use two fMRI datasets (one for encoding, one for decoding), and for each we consider two validation criteria. The first criterion is common to both datasets and assesses whether the learned HRF fits better unseen data than the canonical function. We fit a rank-one regression model on all but one session. This gives us an estimate of the HRF for all voxels. We then use this HRF to compute the log-likelihood of a GLM model on unseen data. In similar fashion, we compute the log-likelihood using the canonical HRF. A paired t-test allows us to conclude whether both likelihoods are significantly different and thus if one model has better goodness of fit.

This criterion serves to validate our model using a likelihood function defined on the raw fMRI timeseries. However, we are mainly interested in the GLM as a pre processing step for higher level analysis such as decoding or encoding models. That is why, in order to fully assess the relevance of this contribution, we propose to quantify the performance of our rank-one model using the performance metrics commonly used with decoding and encoding models.
described in [10], where the task consists in predicting the voxel activity, before being subtracted from 1 (predictive scoring).

Linear predictions on the left out fold were compared to the true activations in 80% (4 of 5 sessions, i.e. 1400 images) of the training set. We used Ridge regression to learn a predictor of voxel activity on 80% (4 of 5 sessions, i.e. 1400 images) of the training set. For the second images, a basis of 6 elements consisting on the canonical HRF and its five successive derivatives was chosen.

A. Data description

Natural images For the encoding model, we use the publicly available dataset from [9], where the task consists in predicting the activation maps from a set of natural images. Subjects viewed 1750 training images, each presented twice, and 120 validation images, each presented 10 times, while fixating a central cross. Images were flashed 3 times per second (200ms on-off-on-off-on) for one second every 4 seconds, leading to a rapid event-related design. The data was acquired in 5 scanner sessions, each comprising 5 blocks of 70 training images, each presented twice within the block and 2 blocks of validation images showing 12 images each 10 times.

Evaluation of the performance of our GLM was done with a simple encoding task: Using a spatially smoothed Gabor pyramid transform modulus with 2 orientations and 4 scales, we used Ridge regression to learn a predictor of voxel activity on 80% (4 of 5 sessions, i.e. 1400 images) of the training set. Linear predictions on the left out fold were compared to the true activations in $l_2$ norm and normalized by the variance of the voxel activity, before being subtracted from 1 (predictive $r_2$ scoring).

Word decoding For the decoding task, we use the dataset described in [10], where the task consist in predicting the visual percept formed by four letter words. Each word was presented on the screen for 3s at a flickering frequency of 15Hz. A 5s rest interval was inserted between each word presentation. The subject was asked to fixate a colored cross at the center of the screen. Each session comprised 46 words including 6 verbs. To ensure that subjects were reading, they were asked to report with a button press when a verb was presented on the screen. Repetitions corresponding to verbs were then removed from the analysis. Six acquisition blocks were recorded, leading to 240 different words used in the analysis. We evaluate the performance of our GLM in this decoding task by calculating the percentage of overall correctly predicted bars forming the image.

B. Results

We first report results on the natural images dataset. As described above, we validate by fitting a rank-one regression model on all but one session. On the left-out session, we found the log-likelihood of the GLM model obtained with the data-driven HRF to be consistently larger than the log-likelihood obtained using a canonical HRF. A paired difference test was used to conclude that the mean likelihood of the rank-one model is significantly larger with p-value $< 10^{-3}$.

The estimated HRF across the 100 most responding voxels is presented in Fig. 2. We show the average value for the learned HRFs across voxels together with its standard deviation. As expected, the data-driven HRFs resemble the canonical HRF, however, the peak is located on average one second before the peak of the canonical function.

The predictive $r_2$ scores of 100 voxels are shown in a scatter plot in Fig. 3. We chose the 100 best predicted voxels using a classic GLM. These scores are on the x-axis. The y-axis shows the scores using our method. Significantly more scores lie above the diagonal ($p < 10^{-4}$), Wilcoxon signed rank test), suggesting that that learning the HRF is beneficial to this encoding scheme.

The second dataset that we analyze is the word-decoding dataset. This dataset is an asynchronous setting with TR = 2.4 s and trials of 8 s. As described above we chose to oversample by a factor or 3 the original TR to a TR of 0.8s. Because of the low time resolution and the shorter time course, we chose to constrain the HRF to a small subspace in order to avoid overfitting associated with complex models. The set of basis functions for this subspace is given by the canonical HRF and its five successive derivatives. We have experimented with several different bases including Chebyshev polynomials, Fourier basis and discrete cosine basis. All these basis capture the general trend of the HRF function and give similar results, but each set induces some bias towards specific shape functions, while this set of generators favors more biologically plausible models. As with the previous dataset, we observed log-likelihood values that are constantly larger than for the GLM. A paired difference test was used to conclude that the mean log-likelihood is significantly higher with p-value $< 3 \times 10^{-3}$.
As can be seen in Fig. 2, second image, the estimated HRF resembles the canonical HRF, which is not surprising given the subspace in which it is constrained. As with the previous dataset, the peak of the HRF is slightly advanced with respect to the canonical function.

We used the activation patterns estimated by the rank-one regression model as input to the decoding study. Within the four subjects analyzed, we observed a systematic increase in the mean score, ranging for +1% to +3% (out of a score of 65%). We observed higher increase for the best performing subjects. A Wilcoxon signed-rank test on the scores across subjects gave us a p-value < 0.07 for the significance of these differences.

IV. CONCLUSION

We have presented a model that jointly estimates the hemodynamic response function (HRF) and the activation patterns from the BOLD signal. This model, named rank-one regression, can be optimized using standard smooth optimization methods such as L-BFGS. We investigated whether this model yields better prediction for encoding and decoding models.

In a first step, we assessed the quality of the HRF estimation by comparing the likelihood of the GLM on unseen data using both the estimated HRF and the canonical HRF. To assess the impact on encoding and decoding models, we have selected two fMRI datasets and used the GLM coefficients obtained by our rank-one model as input data.

We found out that using the activation patterns estimated by the rank-one model significantly improved encoding and decoding studies. In the encoding study we found a generalized improvement across voxels. In the decoding study observed improved scores for all subjects across the study.

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REFERENCES