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Submitted on 25 Mar 2015

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Search for muon neutrinos from gamma-ray bursts with the ANTARES neutrino telescope using 2008 to 2011 data

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Aims. We search for muon neutrinos in coincidence with GRBs with the ANTARES neutrino detector using data from the end of 2007 to 2011. 

Methods. We search for muon neutrinos in coincidence with GRBs with the ANTARES neutrino detector using data from the end of 2007 to 2011.

Results. Expected neutrino fluxes were calculated for each burst individually. The most recent numerical calculations of the spectra using the NeutCosmA code were employed, which include Monte Carlo simulations of the full underlying photodisintegration interaction processes. The discovery probability for a selection of 296 GRBs in the given period was optimised using an extended maximum-likelihood strategy. 

Conclusions. No significant excess over background is found in the data, and 90% confidence level upper limits are placed on the total expected flux according to the model.

Key words. neutrinos – gamma-ray burst: general – methods: numerical

1. Introduction

Gamma-ray bursts (GRBs) are short and very intense flashes of high-energy gamma rays, which occur unpredictably and isotropically over the sky (Meegan et al. 1992). Over timescales of the order of a few seconds, they release as much energy as the Sun in its entire lifetime (see Woosley & Bloom 2006, for a review). The doubly peaked distribution of burst durations measured by the BATSE satellite led Kouveliotou et al. (1993) to the classification of GRBs into two types (see also Paciesas et al. 1999). The sub-class of long bursts (with durations $\lesssim 2$ s) has been shown to be associated with supernovae of type I b/c (see e.g. Galama et al. 1998). For the short bursts ($\lesssim 2$ s duration) it is much harder to measure the fast-fading afterglow emission and thereby obtain information about their origin. However, they are now widely accepted to originate from the merging of two compact objects, for example neutron stars and black holes (Eichler et al. 1989; Nakar 2007).

In the fireball model (as proposed by Mészáros & Rees 1993), the observed electromagnetic radiation is explained by highly relativistic outflows of material, most likely collimated
charged-current interactions of cosmic muon neutrinos with matter in or close-by to the detector. The passage of these muons through the seawater induces the emission of Cherenkov light that is then detected by an array of photo-multiplier tubes (PMTs). Using the time and position information of the photons, the muon trajectory is reconstructed. The original neutrino direction is then inferred from the measured muon direction – at the energies considered here, the uncertainty introduced by the scattering angle is negligible compared to the detector’s resolution.

The ANTARES telescope is located in the Mediterranean Sea at a depth of 2.4 km. The detector consists of twelve vertical “strings” anchored to the seabed, each of which is held upright by a buoy at the top. They are separated from each other by a typical distance of 70 m. The twelve strings, each with a length of 450 m, are equipped with 25 triplets of PMTs, building a three-dimensional array of 885 PMTs in total. The triplets have a vertical spacing of 14.5 m between them, whereas the first triplet is placed at a height of 100 m above the seabed.

Bioluminescence and radioactive decay of $^{40}$K produce a random optical background that can vary between 50 and 300 kHz per PMT, depending for example on the time of the year or the sea current. A multi-level online triggering procedure is applied to select possible particle signatures – see Aguilar et al. (2007) for a more detailed description.

In addition to the cosmic neutrino signal that the ANTARES experiment is searching for, there are other processes that can produce muon tracks in the detector that are considered as background events. Air showers are generated when high-energy cosmic rays hit the Earth’s atmosphere. Among other particles, muons and neutrinos are produced in these showers. Since only the (weakly interacting) neutrinos are capable of traversing the Earth, it can effectively be used as a “shield” against all particles but neutrinos. By searching only for upgoing particles therefore, the atmospheric downgoing muon background can be rejected. Nevertheless, muons from above can also produce signals in the detector that appear as upgoing events. Using quality cuts on the reconstruction parameters, these falsely reconstructed atmospheric muon tracks can be suppressed to a rate of 0.4 events per day.

Atmospheric neutrinos produced by cosmic rays below the horizon can also traverse the Earth, and represent the main background component (~three events per day after quality cuts, see Adrián-Martínez et al. 2012) to the cosmic neutrinos.

In the analysis presented in the following, the requirement of temporal and spatial coincidence with a recorded GRB reduces the number of expected background events to ~$O(10^{-4})$ per GRB (see Sect. 6), and an extended likelihood method is furthermore used to distinguish between signal and background events. Data collected between December 6, 2007 and the end of 2011 were analysed. The first six months of this period comprised the last phase of construction of the apparatus, after the deployment of detection lines 6 to 10, with the last two lines installed in May 2008. In that period, the instrumented volume of the detector increased from 0.008 to 0.011 km$^3$ at full size. The corresponding average effective area to muon neutrinos as a function of the energy is shown in Fig. 1 for different declination bands. The huge increase of effective area with energy, a common feature of neutrino telescopes, is due not only to the increase of the neutrino cross-section, but also to that of the muon range, which can reach several kilometres at the highest

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1 Throughout the paper, “neutrino” will denote both $\nu$ and $\bar{\nu}$, and “muon” will denote both $\mu$ and $\bar{\mu}$.

2 One string is equipped with 60 instead of 75 PMTs.
energies. The total integrated livetime of the data in coincidence with the selected 296 GRB search-time windows is 6.6 h.

3. GRB selection and parameters

The GRB parameters needed for the search and the simulation of expected neutrino fluxes are primarily obtained from different tables provided by the Swift (Gehrels et al. 2004) and Fermi (Atwood et al. 2009; Meegan et al. 2009) collaborations. This information is then supplemented using a table supplied by the IceCube Collaboration (Aguilar 2011), which is created by parsing the Gamma-ray Coordinates Network (GCN) notices1. In Appendix A, we specify how these tables are merged, how often burst parameters are taken from each of them, and how the search-time windows are defined.

For the final sample, GRBs are required to meet certain criteria as specified in Table 1 – short bursts, for instance, are excluded as this class is much less understood. A total of 296 bursts pass these selection cuts, of which 10% are also included in the most recent gamma-ray-burst search from IceCube (Abbasi et al. 2012). The distribution of the selected bursts in equatorial coordinates is shown in Fig. 2. Out of this selection, GRB 110918 outshines all others by at least half an order of magnitude in the expected neutrino flux (see Sect. 4). It is at the same time one of the most intense bursts ever observed by the Konus-Wind instrument (Aptekar et al. 1995; Golenetskii et al. 2011). Unfortunately, both Swift and Fermi satellites were Earth-occulted at the time of the burst (Krimm et al. 2011), but Swift could still observe the afterglow emission after ~30 h. The measured parameters for this exceptional burst are given in Table A.2.

1 GCN: http://gcn.gsfc.nasa.gov/gcn3_archive.html

4. Calculation of neutrino spectra

To calculate the expected neutrino spectra, we focused on the recently developed fully numerical NeuCosmA model (Hümmer et al. 2010, 2012). In addition, we also present the widely used analytical approach of Guetta et al. (2004) in the following.

4.1. Analytic approaches

Waxman & Bahcall (1997) were the first to calculate the expected neutrino flux in coincidence with the electromagnetic GRB in the framework of the standard fireball internal shock model, using averaged burst parameters as measured by the BATSE instrument on board the CGRO satellite (Band et al. 1993). Their calculation was based on the assumption of Fermi-accelerated protons in the relativistic ejecta of the burst interacting with the associated photon field to produce pions via the $\Delta$-resonance. The subsequent decay of charged pions and muons leads to the emission of high-energy neutrinos. The authors derived a doubly broken power-law spectrum for the neutrinos. Their model is referred to as the standard Waxman-Bahcall GRB neutrino flux, and is for instance used to set limits with the BAikal (Avrorin et al. 2011) and AMANDA (Achterberg et al. 2008) experiments. Guetta et al. (2004) modified the formulae of Waxman and Bahcall to calculate individual neutrino fluxes for the bursts. Such individual burst predictions are used in recent searches with IceCube (Abbasi et al. 2010, 2011, 2012), RICE (Besson et al. 2007), and ANITA (Vieregg et al. 2011). Note in particular that the most recent IceCube limit was a factor of 3.7 below predictions made using this model. This could either indicate the need for rejection of the model, a modification of the parameters upon which it is based, or for more detailed modelling of the neutrino emission within the fireball paradigm.

To calculate the analytic spectrum as shown in Fig. 3a, blue solid line, the formulae given in Abbasi et al. (2010, Appendix A) were applied. In principle, Guetta et al. (2004) predicted different break energies for $\nu_\mu$ and $\bar{\nu}_\mu$ (see Guetta et al. 2004, Eqs. (A.10) and (A.11)), yielding three breaks in the combined $\nu_\mu + \bar{\nu}_\mu$ spectrum (or in the single $\nu_\mu$ spectrum when taking oscillations into account) as shown in Fig. 3a, blue dashed – in the previous ANTARES analysis (Adrián-Martínez et al. 2013), this effect has been accounted for.
Fig. 3. a) Expected $\nu_{\mu} + \bar{\nu}_{\mu}$ spectra of GRB 110918. The analytic model of Guetta et al. (2004) (blue) is shown with the usual simple treatment (blue solid) and accounting for different break energies of $\nu_{\mu}$ and $\bar{\nu}_{\mu}$ (blue dashed). The numerical NeuCosmA prediction is presented in red. b) Individual $\nu_{\mu} + \bar{\nu}_{\mu}$ NeuCosmA spectra of the 296 GRBs (thin lines) and their sum (thick line).

4.2. NeuCosmA model

Monte Carlo algorithms such as SOPHIA (see Mücke et al. 2000) have improved the calculations of photohadronic interactions in astrophysical processes. Hümmer et al. (2012) showed that a more detailed treatment of the particle physics involved in calculating the neutrino spectra greatly changes the resulting neutrino flux predictions compared with those described in Sect. 4.1. Using their neutrinos from cosmic accelerators (NeuCosmA) code, the authors pointed out how the full photohadronic interaction cross-section, individual treatment of secondary particles (including energy losses), and neutrino mixing affect the predicted neutrino flux (Hümmer et al. 2010, 2012; Baerwald et al. 2012). Owing to these effects, the doubly peaked structure of the spectrum from muon and pion decays now features an additional high-energy component from $K^+$ decays. Moreover, the authors discussed several problems with the older neutrino flux estimations, such as using the real average photon energy instead of the peak energy of the photon distribution, taking the full width of the $\Delta$-resonance into account, and simulating the energy losses of secondary particles as well as the energy dependence of the mean free path of protons (see also He et al. 2012). The combination of all these effects gives rise to a prediction for the neutrino yield that is about one order of magnitude below the result of Guetta et al. (2004).

Note that the NeuCosmA model, however, does not introduce any new assumptions on the nature of GRBs in general, but applies known physics in greater detail within the paradigm of the fireball model. Hence, limits on these predictions will first of all constrain parameters governing the fireball model (such as the boost factor $\Gamma$ of the jet), and might later on probe the fireball paradigm itself.

Predictions made by the NeuCosmA model are used to optimise this analysis. In Fig. 3a, a comparison between the predicted neutrino spectra from Guetta et al. (2004) and from NeuCosmA is shown exemplarily for GRB 110918. Figure 3b shows all individual numerical neutrino spectra and their sum.

5. Simulation and signal probability density function

For each GRB in the selection, neutrino events were generated with high statistics to simulate the predicted NeuCosmA spectrum. They were then reconstructed to compute the acceptance of the detector. Their distribution gives the signal probability density function (PDF) labelled $S(\alpha) = dN(\alpha)/d\Omega$, with $\alpha$ being the space angle between the reconstructed track direction and the gamma-ray-burst’s coordinates.

Firstly, the passage of neutrinos from the direction of a GRB through the Earth was simulated. If they interacted with matter sufficiently close to or in the detector, the resulting hadronic shower from the break-up of the target nucleon was generated. The secondary particles were then propagated through the medium, inducing the emission of Cherenkov photons on their trajectory. The response of the PMTs to the emitted Cherenkov light was then simulated, taking into account the detector and environmental conditions at the time of the GRB occurrence. For a more detailed description of the simulation scheme, see Adrián-Martínez et al. (2012).

For each GRB, $4 \times 10^5$ muon neutrinos were simulated. No background events were generated, as the background rate was estimated using the data themselves (see Sect. 6).

The applied track reconstruction algorithm is the same as used in Adrián-Martínez et al. (2012). It is based on a multi-step algorithm to maximise the likelihood of an assumed track hypothesis. The reconstruction returns two quality parameters, namely the track-fit quality parameter $\Lambda$ and the estimated angular uncertainty on the muon track direction $\beta$. Cuts on these parameters can be used to improve the signal-to-noise ratio. To ensure a good directional reconstruction of the selected neutrino candidates, we required $\beta < 1^\circ$. For a sample of track candidates with high reconstruction quality ($\Lambda > -5.2$), this cut has been shown to remove most of the remaining atmospheric muons that were falsely reconstructed as upgoing without significantly affecting the neutrino signal: atmospheric muons are reduced by a factor of $1.2 \times 10^{-2}$, while atmospheric neutrinos are reduced by 0.19. The total background due to atmospheric events is decreased by $10^{-4}$ (Adrián-Martínez et al. 2012). In this analysis, the same cut combination would leave $-60 \ldots 70\%$ of a typical GRB signal. However, the narrow time windows (typically a few tens of seconds, see Table 2 for per-GRB values) yield intrinsically low background in coincidence with each GRB, allowing one to loosen the cuts on the $\Lambda$ parameter. These were then optimised for each burst individually.

To account for the satellite’s uncertainty on the direction of the GRB, the reconstructed space angle was additionally smeared with a Gaussian of the appropriate width.

The resulting distribution of events relative to the GRB direction for each cut on $\Lambda$ was then fitted with the function

$$
\log S(\alpha) = \log \left( \frac{dN(\alpha)}{d\Omega} \right) = \begin{cases} A, & \text{if } \alpha \leq \alpha_0 \\ A - B \cdot \left( 1 - \exp \left( \frac{-\log(\alpha) - \log(\alpha_0)}{\Delta \alpha} \right) \right), & \text{if } \alpha > \alpha_0, \end{cases}
$$

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Table 2. Optimisation results for the ten most promising GRBs.

<table>
<thead>
<tr>
<th>GRB</th>
<th>$\Lambda_{\text{cut}}$</th>
<th>$\mu_0$</th>
<th>$\mu_0^{\text{NeuCosmA}}$</th>
<th>$\mu_{\text{cut}}$</th>
<th>$\langle \alpha \rangle$</th>
<th>$T_{\text{search}}$</th>
<th>$\sigma_{\text{cut}}$</th>
</tr>
</thead>
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<tr>
<td>11091889</td>
<td>-5.3</td>
<td>$1.1 \times 10^{-4}$</td>
<td>$3.0 \times 10^{-2}$</td>
<td>$1.5 \times 10^{-1}$</td>
<td>0.30</td>
<td>73.4</td>
<td>5$\sigma$</td>
</tr>
<tr>
<td>11091889</td>
<td>-5.5</td>
<td>$3.7 \times 10^{-4}$</td>
<td>$3.5 \times 10^{-2}$</td>
<td>$1.7 \times 10^{-1}$</td>
<td>0.32</td>
<td>73.4</td>
<td>5$\sigma$</td>
</tr>
<tr>
<td>08060725</td>
<td>-5.4</td>
<td>$5.5 \times 10^{-4}$</td>
<td>$6.5 \times 10^{-3}$</td>
<td>$1.4 \times 10^{-2}$</td>
<td>0.33</td>
<td>164.3</td>
<td>5$\sigma$</td>
</tr>
<tr>
<td>11100892</td>
<td>-5.5</td>
<td>$3.6 \times 10^{-4}$</td>
<td>$2.2 \times 10^{-3}$</td>
<td>$2.6 \times 10^{-2}$</td>
<td>0.35</td>
<td>75.4</td>
<td>5$\sigma$</td>
</tr>
<tr>
<td>10101417</td>
<td>-5.1</td>
<td>$4.1 \times 10^{-4}$</td>
<td>$1.2 \times 10^{-3}$</td>
<td>$1.7 \times 10^{-2}$</td>
<td>0.89</td>
<td>723.1</td>
<td>5$\sigma$</td>
</tr>
<tr>
<td>10072809</td>
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<td>$2.0 \times 10^{-4}$</td>
<td>$9.6 \times 10^{-4}$</td>
<td>$1.4 \times 10^{-3}$</td>
<td>0.49</td>
<td>268.6</td>
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</tr>
<tr>
<td>09020174</td>
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<td>$7.0 \times 10^{-4}$</td>
<td>$2.4 \times 10^{-3}$</td>
<td>0.39</td>
<td>126.6</td>
<td>5$\sigma$</td>
</tr>
<tr>
<td>11122048</td>
<td>-5.2</td>
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<td>$6.2 \times 10^{-4}$</td>
<td>$1.2 \times 10^{-3}$</td>
<td>1.13</td>
<td>66.5</td>
<td>5$\sigma$</td>
</tr>
<tr>
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<td>-5.4</td>
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<td>$3.9 \times 10^{-4}$</td>
<td>$5.7 \times 10^{-3}$</td>
<td>1.02</td>
<td>112.1</td>
<td>5$\sigma$</td>
</tr>
<tr>
<td>11062215</td>
<td>-5.4</td>
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<td>$4.3 \times 10^{-4}$</td>
<td>$9.5 \times 10^{-3}$</td>
<td>1.42</td>
<td>116.6</td>
<td>5$\sigma$</td>
</tr>
<tr>
<td>08100914</td>
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<td>$1.3 \times 10^{-4}$</td>
<td>$3.5 \times 10^{-4}$</td>
<td>$1.9 \times 10^{-3}$</td>
<td>0.94</td>
<td>70.2</td>
<td>3$\sigma$</td>
</tr>
<tr>
<td>all GRBs:</td>
<td>mean</td>
<td>$1.7 \times 10^{-4}$</td>
<td>$2.0 \times 10^{-4}$</td>
<td>$1.6 \times 10^{-3}$</td>
<td>2.85</td>
<td>80.4</td>
<td></td>
</tr>
<tr>
<td>sum</td>
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<td>$6.1 \times 10^{-2}$</td>
<td>$4.8 \times 10^{-1}$</td>
<td>$2.4 \times 10^{0}$</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Notes. Optimised $\Lambda_{\text{cut}}$ values for the ten GRBs with the highest discovery probabilities and the resulting expected number of background and signal events $\mu_0$ and $\mu_{\text{cut}}$ at the significance level $\sigma_{\text{cut}}$. The consequent median angular spread of events $\langle \alpha \rangle$ is also provided. In the last rows, the sum and mean of the values for all 296 GRBs at 3$\sigma$ is given. The naming convention for the GRBs is similar to that used by Fermi, the last two digits of the GRB name correspond to the fraction of the day at which the burst occurred. The full table is available at the CDS.

Fig. 4. Simulated and reconstructed signal events per solid angle $\Omega$ versus the logarithm of the space angle $\alpha$ in degrees for the burst GRB 110918: muon tracks are plotted in black, shower-like events are drawn in green. The corresponding fit is shown in red (see Eq. (1)). The grey dotted line indicates the median angular spread of events $\langle \alpha \rangle = 0.32^\circ$. The blue dashed line shows the flat background distribution $B(\alpha)$ as calculated in Sect. 6. Cut values $\Lambda > -5.5$ and $\beta < 1^\circ$ are applied here.

6. Background estimation

Upgoing atmospheric neutrinos constitute the main background component for each GRB, with a smaller contribution coming from misreconstructed downgoing atmospheric muons. To estimate the mean number of background events $\mu_0$ for each $\Lambda$ cut for each burst as realistically as possible, data were used. However, as the number of upgoing events is very low (~four per day, see Adrián-Martínez et al. 2012, Table 1), long time periods are needed to yield sufficient statistics, which in turn requires averaging over different data-taking conditions (in particular because of seasonal variations of the optical background). To compensate for this, we first estimated the average rate of reconstructed events in the GRB’s direction $\langle n(\theta, \phi)_{\text{GRB}} \rangle$ all runs in local coordinates zenith $\theta$ and azimuth $\phi$ using data from the entire 2007 to 2011 period, then adjusted it for varying data-taking conditions (for a detailed description, see Appendix B).

For the final search, the number of background events in coincidence with each burst and within a search cone of 10$^\circ$ around its position was evaluated to be of the order of 10$^{-3}$. The background PDF, $B(\alpha)$, was considered to be flat in $\Omega$ within this cone, as shown in Fig. 4 for GRB 110918.

7. Pseudo-experiments and extended maximum-likelihood ratio

In the following, we show how pseudo-experiments were generated to compute the significance of a measurement, then how the cut on the reconstruction quality parameter $\Lambda$ was optimised for each GRB to yield the highest discovery probability for a signal according to the NeuCosmA model.

For the pseudo-experiments, signal and background events $i$ with space angle $\alpha_i$ were drawn randomly from the normalised signal $S(\alpha)$ and background $B(\alpha)$ PDFs corresponding to each chosen cut on $\Lambda$. For each pseudo-experiment with a total number of events $n_{\text{tot}}$, the test statistic $Q$ was calculated as follows:

$$Q = \max_{\mu_0, \mu_{\text{cut}}^{\text{B}}} \left\{ \sum_{i=1}^{n_{\text{tot}}} \log \left( \mu_0^i S(\alpha_i) + \mu_{\text{cut}}^{iB} B(\alpha_i) \right) / \mu_{\text{cut}}^{iB} B(\alpha_i) - (\mu_0^i + \mu_{\text{cut}}^{iB}) \right\}. \quad (2)$$

This is the so-called extended maximum-likelihood ratio (Barlow 1990) with an a priori knowledge of the expected number of background events $\mu_{\text{cut}}^{iB}$ (as evaluated in Sect. 6 and Appendix B). Higher values of $Q$ indicate that the measurement is more compatible with the signal hypothesis. The signal contribution $\mu_0^i$ is scanned between 0 and $n_{\text{tot}}$, its value corresponding to the maximum of the sum in Eq. (2) returns the estimated signal $\mu_0^i$.

In the following, $h_n(Q)$ denotes the distribution of $Q$-values for $n_i$ injected signal events with a Poisson-distributed number of background events with expectation value $\mu_{\text{cut}}^{iB}$. The significance of a measurement is determined by its $p$-value$^4$, which is given by the probability to yield $Q$-values at least as high as that observed if the background-only hypothesis were true. Hence, using the background-only distribution $h_n(Q)$, the lowest $Q$-value $Q_p^{\text{thres}}$ that is necessary to claim a discovery with a certain $p$-value can be calculated with

$$P(Q \geq Q_p^{\text{thres}} | \mu_0) = \int_{Q_p^{\text{thres}}}^{\infty} h_n(Q) dQ = p. \quad (3)$$

The probability distributions $h_n(Q)$ for different $n_i$ are shown in Fig. 5, with the threshold $Q$-values indicated by the grey dashed

$^4$ We used the two-sided convention, that is, $p_{2\sigma} = 2.7 \times 10^{-3}$, $p_{3\sigma} = 5.7 \times 10^{-7}$. 

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The probability distribution of $Q$ values for any number of expected signal events $\mu_s$ is calculated via

$$P(Q | \mu_s) = \sum_{n_s = 0}^{\infty} P(n_s | \mu_s) h_{th}(Q),$$

(4)

with the Poisson distribution $P(n_s | \mu_s)$ giving the probability of observing $n_s$ events from a mean number of expected events $\mu_s$. The integral of $P(Q | \mu_s)$ gives the model discovery potential $MDP$; it is the probability to make a discovery assuming that the model was correct:

$$MDP = \int_{Q_{thres}}^{\infty} P(Q | \mu_s) dQ$$

(5)

$$= \sum_{n_s = 0}^{\infty} \left( \prod_{i=1}^{N_{GRB}} P(Q(\mu_i) \mid h_{th}(Q)) \right) dQ,$$

The value of the $\Lambda$ cut at a given significance level is then chosen as that which maximises the $MDP$ for the value of $\mu_s$ predicted by the NeuCosmA model (see Sect. 4.2). Figure 6 shows $MDP(\mu_s)$ of GRB 110918 for $3\sigma$, $4\sigma$, and $5\sigma$ versus an arbitrary number of signal events. The distribution $P(Q | \mu_s)$ from Eq. (4) was also used to set upper limits on the number of signal events when no discovery is made. For example, when the final analysis returns $Q_{\text{mean}}$, set a 90% confidence level (C.L.) upper limit $\mu_s^{90\%}$ on the signal by rejecting all event expectations $\mu_s$ that would lead to values $Q > Q_{\text{mean}}$ in 90% of all pseudo-experiments:

$$P(Q \geq Q_{\text{mean}} | \mu_s^{90\%}) = \int_{Q_{\text{mean}}}^{\infty} P(Q | \mu_s^{90\%}) dQ = 0.9.\quad (6)$$

When no event was measured ($Q_{\text{mean}} = 0$), a 90% C.L. upper limit was set at 2.3, the lowest possible value.$^5$

For each number of injected signal events $n_s$, $10^5$ pseudo-experiments were generated$^6$ to derive the signal distributions $h_{th}(Q)$. Using this procedure, the model discovery potential was calculated for any given $\mu_s$, and the final cut on $\Lambda$ for each GRB was found.

$^5$ The value derives from Poisson statistics, since the probability to detect at least one event at a mean rate of 2.3 is exactly 90%.

$^6$ Far more background-only pseudo-experiments are required to allow determining $Q_{\text{thres}}$ at $p$-values as low as $p_{3\sigma}/296 \sim 2 \times 10^{-4}$.

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8. Search optimisation

In the following, the best trade-off between an increased sample and the associated statistical penalty is investigated. In general, a weighting factor, $w_i$, could be assigned to each GRB according to the predicted flux from the model. However, this would result in the search being very sensitive to the combined uncertainty from the NeuCosmA model and especially to the parameters upon which it is based. An alternative approach is to include only the $N_{\text{GRB}}$ most promising candidates, where to maintain the same overall probability of making a false discovery, the $p$-value for each burst must be divided by the total trial factor $N_{\text{GRB}}$. By ordering the bursts from highest to lowest $MDP$, the most promising $N_{\text{GRB}}$ can be chosen to maximise the combined probability of making a significant discovery from any of them. The total model discovery potential is then calculated via

$$MDP(N_{\text{GRB}}) = 1 - \prod_{i=1}^{N_{\text{GRB}}} (1 - MDP_i),\quad (7)$$

The resulting distributions of $MDP(N_{\text{GRB}})$ as a function of the size of the considered sub-sample of GRBs are shown in Fig. 7 for $3\sigma$, $4\sigma$, and $5\sigma$ (thick lines). The $3\sigma$ $MDP$ distribution rises to a maximum of 0.059 at a sample size of 106 GRBs, but is relatively flat around its highest value. For a search optimised for $3\sigma$, it is therefore reasonable to take the whole set of GRBs, as it does not decrease the $MDP$ significantly (around 3.4%) and the search remains less model-dependent. The $5\sigma$ $MDP$ distribution, on the other hand, is prominently peaked at $N_{\text{GRB}} = 1$, with $N_{\text{GRB}} = 2$ being almost equivalent ($MDP(1) = MDP(2) = 0.025$) — the model discovery potential then decreases for larger samples. Even for $N_{\text{GRB}} = 2$, the second strongest GRB contributes only a small fraction to the discovery potential.

For comparison, the distributions for a simple counting search (in which all events passing the cut criteria carry equal weighting) are also shown in Fig. 7 (thin lines). To mimic this search, a radius cut $\alpha_{\text{cut}}$ for each GRB was calculated from the known background $\mu_b$ at fixed reasonable quality cuts ($\Lambda > -5.5, \beta < 1^\circ$) and the given significance level $p/N_{\text{GRB}}$. Applying this search radius cut on the signal PDF $S(\alpha)$, the expected number of signal events $\mu_s$ can be estimated and consequently the $MDP$ evaluated as the probability of detecting more than zero events.

As expected, the $MDP$ curves for a counting analysis are well below those for the likelihood method, showing the advantage of the search method used in this analysis. The shapes of the
curves, on the other hand, are quite similar, and the same conclusions can be drawn from them, namely that using the whole sample gives the best discovery probability at 3σ and using only the individual GRB 110918 at 5σ.

Based on these results, we decided to optimise the quality cut on λ for a likelihood search on the whole sample of 296 GRBs at the 3σ significance level. Because 3σ is not enough to claim a discovery, we predefined a cut on λ that was optimised for a 5σ discovery, which was then used for a separate search for the emission from GRB 110918 only.

Optimised cuts λ_{cut} for the final analysis as well as the accordingly expected number of background and signal events, the median angular resolution, and the search-time window are shown in Table 2 for the ten most promising GRBs. A full list for the 296 selected bursts can be found at the CDS.

9. Results and discussion

Using the strategy outlined above, we analysed ANTARES data from the end of 2007 to 2011 searching for neutrino events in coincidence with the search-time windows and within 10° around each GRB. No data events passed this selection within the accumulated search duration of 6.6 h. Hence, the measured Q-value is zero.

In total, 0.06 neutrino events from GRBs are expected from the NeuCosmA model, where only a small contribution of 4.6 × 10^{-5} events is not due to particle tracks produced by muon neutrinos – the Guetta model predicts 0.5 signals from muon neutrinos. The overall background in the 10° cones is 0.05 events. The 90% C.L. upper limits on the expected number of signal events μ_s from each model are thus set to 2.3 events, and the corresponding limits on the muon neutrino flux, F_ν^μ, from GRB 110918 as well as on the cumulative flux from the whole sample are shown in Fig. 8. The simple treatment of the Guetta model is represented here (see Fig. 3a, solid line). For the NeuCosmA model, the limit on the total flux lies a factor of 38 above the expected spectrum (4.4 for Guetta). The right-hand axis of Fig. 8b represents the limits translated into limits on the inferred quasi-diffuse neutrino flux:

\[ E^2 \Phi_\nu = \sum E^2 F_\nu \times \frac{1}{4\pi N_{GRB}} 667 \text{ y}^{-1}, \quad (8) \]

where \( v = v_\nu + \bar{v}_\mu \), assuming that each analysed sample represents an average burst distribution and that the annual rate of long bursts is 667 per year.

The first ANTARES limit (Adrián-Martínez et al. 2013) obtained for 40 GRBs during the construction phase of the detector in the year 2007 is also shown in Fig. 8b. That analysis was based on the Guetta model (accounting for different break energies of \( v_\nu \) and \( \bar{v}_\mu \)) and employed a counting method searching for neutrino events in a two-degree cone around each burst. Using the data from the IC 40 and IC 59 detector phases in 2008 to 2010, IceCube recently published a more stringent limit on the neutrino emission as predicted by the “simple” Guetta model (Abbasi et al. 2012), which is also shown in Fig. 8b.

Because of the larger effective area of the IceCube detector, the new ANTARES limit presented in this paper does not set additional constraints on the Guetta emission model. Note, however, that both detectors have complementary sky coverage and therefore the analysed sample of GRBs differs significantly. 90% of the analysed bursts have not previously had their neutrino emission constrained. When comparing limits obtained in different analyses, however, one should keep in mind that the precise shapes of the spectra – and thus, of the limits – depend on the actual selected sample, the measured parameters of the individual bursts and their uncertainty, the set of default parameters and on the chosen model.
10. Conclusion

Using data from the ANTARES detector, a search for muon neutrinos in coincidence with 296 GRBs occurring between the end of 2007 and 2011 has been performed. No events passed the selection criteria and limits on the neutrino flux were derived. For the NeuCosmA model, a limit on the neutrino flux has been shown to be an order of magnitude below that predicted by previous analytic approaches. This helps to resolve the tension between the non-observation of a neutrino signal and the most stringent experimental constraint currently available (Abbasi et al. 2012), which was a factor of 3.7 below the predictions made by the Guetta model.

Hence, existing limits do not yet constrain realistic neutrino emission models based on an internal shock scenario. Nevertheless, the collection of more data with active experiments such as ANTARES and IceCube, as well as with the planned neutrino telescope KM3NeT, will certainly allow the widely established fireball paradigm for GRBs to be probed in the near future.

Acknowledgements. The authors acknowledge the financial support of the funding agencies: Centre National de la Recherche Scientifique (CNRS), Commissariat à l’énergie atomique et aux énergies alternatives (CEA), Agence National de la Recherche (ANR), Commission Européenne (FEDER fund and Marie Curie Program), Région Alsace (contrat CPER), Région Provence-Alpes-Côte d’Azur, Département du Var et Ville de La Seyne-sur-Mer, France; Bundesministerium für Bildung und Forschung (BMBF), Germany; Istituto Nazionale di Fisica Nucleare (INFN), Italy; Stichting voor Fundamenteel Onderzoek der Materie (FOM), Nederlandse organisatie voor Wetenschappelijk Onderzoek (NWO), the Netherlands; Council of the President of the Russian Federation for young scientists and leading scientific schools supporting grants, Russia; National Authority for Scientific Research (ANCS), Romania; Ministerio de Ciencia e Innovación (MICINN), Prometeo of Generalitat Valenciana and MultiDark, Spain; Agence de l’Oranent and CNRST, Morocco. We also acknowledge the technical support of Irfferem, AID and Foseleve Marine for the sea operation and the CC-IN2P3 for the computing facilities. We would like to thank Walter Winter for helpful discussions and making it possible to use the NeuCosmA model.

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Table A.1. Usage of the GRB parameter catalogues.

<table>
<thead>
<tr>
<th>Source</th>
<th>Position</th>
<th>Time</th>
<th>Fluence</th>
<th>Spectrum</th>
<th>Duration</th>
<th>Redshift</th>
<th>Start &amp; Stop</th>
</tr>
</thead>
<tbody>
<tr>
<td>Swift/XRT</td>
<td>17.2% [2]</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Swift/UVOT</td>
<td>11.2% [1]</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

**Notes.** The numbers in square brackets give the assigned priority of this source of information with respect to the parameter(s).

Table A.2. Standard gamma-ray-burst parameters as described in the text.

\[
\begin{align*}
\alpha &= 1 \\
\beta &= \alpha + 1 \\
\epsilon_{\text{peak}} &= 200 \text{ keV} \\
z &= 2.15 \\
L_{\text{iso}} &= 10^{52} \text{ erg s}^{-1} \\
\Gamma &= 316 \\
\epsilon_{\gamma} &= 0.1 \\
\epsilon_{\text{B}} &= 0.1 \\
f_c &= 0.1 \\
(\lambda_{\nu_{\gamma \rightarrow \gamma}}) &= 0.2 \\
t_{\text{esc}} &= 0.01 \text{ s}
\end{align*}
\]

Appendix A: GRB selection

The table of the Swift satellite\(^7\) contains data from the three onboard instruments BAT (gamma rays), XRT (X-rays), and UVOT (ultraviolet), ordered with increasing position-magnitude accuracy from arcminutes to sub-arcseconds. The information provides BAT spectral measurements in the energy range from 15 to 150 keV. The Swift BAT2 Catalogue (Sakamoto et al. 2011)\(^8\) provides re-analysed Swift data, so the spectral information therein is considered to be more accurate. The Fermi GBM Burst Catalogue\(^9\) (Goldstein et al. 2012; Paciesas et al. 2012) supplies the best spectral information in the energy range from 10 keV to 1 MeV: the peak flux spectrum and the spectrum averaged over the burst duration (which is eventually used for the neutrino spectrum calculation) is fitted with four different spectral functions. The angular resolution is of the order of degrees. The IceCube Collaboration also provides a table with GRB parameters\(^10\) (Aguilar 2011), which is created by parsing the Gamma-Ray Coordinates Network (GCN) notices\(^11\). This table is used to fill up missing parameter values for GRBs that have been found in at least one of the other tables.

When merging the information on the GRB parameters, we assigned priorities to the measured values according to their considered accuracy. The priorities of parameters are shown in Table A.1 in square brackets, as well as the percentage of how often information was obtained from each source.

When a parameter could not be measured, standard values as given in Table A.2 were used to calculate the spectra. The form of the photon spectrum is determined by the spectral indices \(\alpha\) and \(\beta\) with \(\beta = \alpha + 1\) and the break energy \(\epsilon_{\text{peak}}\) giving their transition. The isotropic luminosity \(L_{\text{iso}}\) can be calculated from the redshift \(z\) and the total measured fluence in gamma rays \(F\) (given in the energy range from \(E_{\text{min}}\) to \(E_{\text{max}}\)) via \(L_{\text{iso}} \approx 4\pi d_L F / L_{\text{iso}}\) with the luminosity distance \(d_L\). In case of unknown redshift \(z\), we took the default value of \(L_{\text{iso}}\). \(T_{90}\) is the time in which 90\% of the fluence is emitted. The other parameters such as the jet Lorentz boost factor \(\Gamma\), the fraction of jet energy in electrons \(\epsilon_{\gamma}\) and in the magnetic field \(\epsilon_{B}\), the ratio of energy in electrons and protons \(f_c\), the average fraction of proton energy transferred to a pion \((\lambda_{\nu_{\gamma \rightarrow \pi}})\) and the variability of the gamma-ray light curve \(t_{\text{esc}}\) are not present in the tables and hence were taken as default. The standard values are the same as given in Aguilar (2011), with some differences to the IC 22 (Abbasi et al. 2010) default values: \(z = 2.0, \Gamma = 300, L_{\text{iso}} = 10^{51} \text{ erg s}^{-1}\). Baerwald et al. (2012) give a very elaborate overview about the NeuCosmA spectra changing with the input parameters.

The parameters of the extraordinarily strong GRB 110918 are presented in Table A.3.

The time window of the search, \(T_{\text{search}}\), for emission from each burst is delineated by the start and stop times as measured by the satellites or, when these are not provided in the catalogues, as \(T_{90} \pm 30\%\). Additionally, we accounted for the detector’s data acquisition uncertainty (0.4 s), the satellite time given in integer seconds (1 s), and the light propagation from a satellite’s data acquisition uncertainty (0 s). The variability of the gamma-ray light curve \(t_{\text{esc}}\) is taken as \(0.01 \text{ s}\) at search-time window.

Table A.3. Gamma-ray-burst parameters of GRB 110918 as described in the text.

\[
\begin{align*}
\alpha &= 1.2 \\
\beta &= 2.0 \\
\epsilon_{\text{peak}} &= 150 \text{ keV} \\
f &= 7.5 \times 10^{-4} \text{ erg cm}^{-2} \\
E_{\text{min}} &= 0.02 \text{ MeV} \\
E_{\text{max}} &= 10 \text{ MeV} \\
\text{UT} &= 21 : 26 : 57 \text{ } \text{dec} = 32.5^\circ \\
\text{RA} &= -27.1^\circ \\
\Delta_{\text{inn}} &= 0.5^\circ \\
T_{90} &= 69.4 \text{ s} \\
\delta_{\text{z}} &= 0.982 \text{ s}
\end{align*}
\]

**Notes.** All values are read from the IceCube table. Values measured by Konus-Wind (Golenetskii et al. 2011), if not marked otherwise, \(^{(a)}\) Measured by the Isaac Newton Telescope (Tanvir et al. 2011), \(^{(b)}\) Integration time of Konus-Wind (Golenetskii et al. 2011), \(^{(c)}\) Determined from Gemini-N (Levan et al. 2011) and the GTC telescope (de Ugarte Postigo et al. 2011).

Appendix B: Background estimation

We estimated the background event rate for each GRB separately. First, the time-averaged reconstructed event rate in the search-time window of Konus-Wind (Golenetskii et al. 2011), if not marked otherwise, \(^{(a)}\) Measured by the Isaac Newton Telescope (Tanvir et al. 2011), \(^{(b)}\) Determined from Gemini-N (Levan et al. 2011) and the GTC telescope (de Ugarte Postigo et al. 2011).

![GRB parameters table](https://swift.gsfc.nasa.gov/docs/swift/archive/grb_table.html)

![GRB selection notes](http://vizier.u-strasbg.fr/viz-bin/VizieR?-source=J/ApJS/195/2)

![GRB catalogues](http://heasarc.gsfc.nasa.gov/W3Browse/fermi/fermigbrst.html)

![GCN notice](http://grbweb.icecube.wisc.edu)

![Standard gamma-ray-burst parameters](http://vizier.u-strasbg.fr/viz-bin/VizieR?-source=J/ApJS/195/2)
calculated by the ratio of the total number of events (in all directions) in the corresponding run $n_i$ to the average total number of events for the respective run duration $t_i$ (see Eq. (B.1)). As $n_i$ may be very small for short runs, the 90\% C.L. upper limit was used instead. Additionally, factors for specific run periods $c_{\text{period}}$ were applied taking into account differences between longer phases of similar run conditions. These values were obtained by fitting the background rate in certain periods separately. This approach assumes that the total number of events – dominated mostly by downgoing atmospheric muons – is proportional to the number of upgoing events. To test this assumption, we determined the measured and estimated rates of upgoing events in longer time periods of a few days, excluding data-taking runs in which GRBs occurred. The measured rate was always found to be $\mu_{\text{meas}} < 1.5 \mu_{\text{est}}$, thus we conservatively increased the estimate by 50\%. Consequently, the expected number of background events in coincidence with each GRB search-time window $T_{\text{search}}$ was calculated via

$$\mu_{\theta,\phi}^{(\theta,\phi)}_{\text{GRB}} = T_{\text{search}} \times \langle n(\theta,\phi)_{\text{GRB}} \rangle_{\text{all runs}} c_{\text{period}} 1.5$$

with $c_j = \frac{\langle n_j \rangle_{\text{all runs}}}{\sum n_j / \sum t_j}$, where $j$ includes all data-taking runs.